# Package 'robustX'

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Type Package
Title 'eXtra' / 'eXperimental' Functionality for Robust Statistics
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<b>Description</b> Robustness 'eXperimental', 'eXtraneous', or 'eXtraordinary'  Functionality for Robust Statistics. Hence methods which are not well established, often related to methods in package 'robustbase'. Amazingly, 'BACON()', originally by Billor, Hadi, and Velleman (2000) <doi:10.1016 s0167-9473(99)00101-2=""> has become established in places. The ``barrow wheel'' `rbwheel()` is from Stahel and Mächler (2009) <doi:10.1111 j.1467-9868.2009.00706.x="">.</doi:10.1111></doi:10.1016>
<b>Imports</b> grDevices, graphics, stats, utils, robustbase (>= 0.92-3)
Suggests MASS, lattice, pcaPP
Enhances ICS
License GPL (>= 2)
Encoding UTF-8
NeedsCompilation no
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# Description

The package **robustX** aims to be a collection of R functionality for robust statistics of methods and ideas that are considered as proposals, experimental, for experiences or just too much specialized to be part of the "Robust Basics" package **robustbase**.

#### **Details**

Package: robustX Type: Package

Title: 'eXtra' / 'eXperimental' Functionality for Robust Statistics

Version: 1.2-6 Date: 2023-01-04

Authors@R: c(person("Martin", "Maechler", role=c("aut", "cre"), email="maechler@stat.math.ethz.ch", comment = c(ORCI

Maintainer: Martin Maechler <maechler@stat.math.ethz.ch>

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Author: Martin Maechler [aut, cre] (<a href="https://orcid.org/0000-0002-8685-9910">https://orcid.org/0000-0002-8685-9910</a>), Werner A. Stahel [aut], Rolf Turner [c

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```

## Author(s)

Werner Stahel, Martin Maechler and potentially others

Maintainer: Martin Maechler

#### See Also

Package **robustbase** which it complements and on which it depends; further package **robust** and the whole CRAN task view on robust statistics, <a href="https://cran.r-project.org/view=Robust">https://cran.r-project.org/view=Robust</a>

## **Examples**

```
pairs( rbwheel(100, 4) )
```

**BACON** 

BACON for Regression or Multivariate Covariance Estimation

# **Description**

BACON, short for 'Blocked Adaptive Computationally-Efficient Outlier Nominators', is a somewhat robust algorithm (set), with an implementation for regression or multivariate covariance estimation.

BACON() applies the multivariate (covariance estimation) algorithm, using mvBACON(x) in any case, and when y is not NULL adds a regression iteration phase, using the auxiliary .lmBACON() function.

#### **Usage**

```
BACON(x, y = NULL, intercept = TRUE,
    m = min(collect * p, n * 0.5),
    init.sel = c("Mahalanobis", "dUniMedian", "random", "manual", "V2"),
    man.sel, init.fraction = 0, collect = 4,
    alpha = 0.05, alphaLM = alpha, maxsteps = 100, verbose = TRUE)

## *Auxiliary* function:
.lmBACON(x, y, intercept = TRUE,
    init.dis, init.fraction = 0, collect = 4,
    alpha = 0.05, maxsteps = 100, verbose = TRUE)
```

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## **Arguments**

Х	a multivariate matrix of dimension $[n \ x \ p]$ considered as containing no missing values.
У	the response (n vector) in the case of regression, or NULL for the multivariate case, where just mvBACON() is returned.
intercept	logical indicating if an intercept has to be used for the regression.
m	integer in 1:n specifying the size of the initial basic subset; used only when init.sel is not "manual"; see mvBACON.
init.sel	character string, specifying the initial selection mode; see mvBACON.
man.sel	only when init.sel == "manual", the indices of observations determining the initial basic subset (and $m <- length(man.sel)$ ).
init.dis	the distances of the x matrix used for the initial subset determined by mvBACON.
init.fraction	if this parameter is $> 0$ then the tedious steps of selecting the initial subset are skipped and an initial subset of size n * init.fraction is chosen (with smallest dis)
collect	numeric factor chosen by the user to define the size of the initial subset (p $^{\ast}$ collect)
alpha	number in $(0,1)$ determining the cutoff value for the Mahalanobis distances (multivariate outlier nomination in mvBACON()), or the discrepancies for regression, see alphaLM.
alphaLM	number in $(0,1)$ where a 1-alphaM t-quantile is the cutoff for the discrepancies (for regression, .lmBACON()); see details.
maxsteps	the maximal number of iteration steps (to prevent infinite loops)
verbose	logical indicating if messages are printed which trace progress of the algorithm.

# **Details**

Notably about the initial selection mode, init.sel, see its description in the mvBACON arguments list.

The choice of alpha and alphaLM:

- Multivariate outlier nomination: see the Details section of myBACON.
- Regression: Let  $t_r(\alpha)$  denote the  $1-\alpha$  quantile of the Student t-distribution with r degrees of freedom, where r is the number of elements in the current subset; e.g.,  $t_r(0.05)$  is the 0.95 quantile. Following Billor et al. (2000), the cutoff value for the discrepancies is defined as  $t_r(\alpha/(2r+2))$ , and they use  $\alpha=0.05$ . Note that this is argument alphaLM (defualting to alpha) for BACON().

# Value

```
BACON(x, y, ...) (for regression) returns a list with components
```

subset the observation indices (in 1:n) denoting a subset of "good" supposedly outlier-free observations.

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tis	the $t_i(y_m, X_m)$ of eq (6) in the reference; the clean "basic subset" in the algorithm is defined the observations $i$ with the smallest $ t_i $ , and the $t_i$ can be
	regarded as scaled predicted errors.
mv.dis	the (final) discrepancies or distances of mvBACON().
mv.subset	the "good" subset from mvBACON(), used to start the regression iterations.

#### Note

"BACON" was also chosen in honor of Francis Bacon:

Whoever knows the ways of Nature will more easily notice her deviations; and, on the other hand, whoever knows her deviations will more accurately describe her ways.

Francis Bacon (1620), Novum Organum II 29.

## Author(s)

Ueli Oetliker, Swiss Federal Statistical Office, for S-plus 5.1; 25.05.2001; modified six times till 17.6.2001.

Port to R, testing etc, by Martin Maechler. Daniel Weeks (at pitt.edu) proposed a fix to a long standing buglet in GiveTis() computing the  $t_i$ , which was further improved Maechler, for **robustX** version 1.2-3 (Feb. 2019).

Correction of alpha default, from 0.95 to 0.05, by Tobias Schoch, see mvBACON.

#### References

Billor, N., Hadi, A. S., and Velleman, P. F. (2000). BACON: Blocked Adaptive Computationally-Efficient Outlier Nominators; *Computational Statistics and Data Analysis* **34**, 279–298. doi:10.1016/S01679473(99)001012

#### See Also

mvBACON, the multivariate version of the BACON algorithm.

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```
(RlmST <- lmrob(log.light ~ log.Te, data = starsCYG))
abline(RlmST, col = "blue")</pre>
```

covNNC

Robust Covariance Estimation via Nearest Neighbor Cleaning

# **Description**

covNNC() estimates robust covariance/dispersion matrices by the nearest neighbor variance estimation (NNVE) or (rather) "Nearest Neighbor Cleaning" (NNC) method of Wang and Raftery (2002, *JASA*).

# Usage

```
covNNC(X, k = min(12, n - 1), pnoise = 0.05, emconv = 0.001, bound = 1.5, extension = TRUE, devsm = 0.01)
```

## **Arguments**

X matrix in which each row represents an observation or point and each column

represents a variable.

k desired number of nearest neighbors (default is 12)

pnoise percent of added noise

emconv convergence tolerance for EM

bound value used to identify surges in variance caused by outliers wrongly included as

signal points (bound = 1.5 means a 50 percent increase)

extension whether or not to continue after reaching the last chi-square distance. The de-

fault is to continue, which is indicated by setting extension = TRUE.

devsm when extension = TRUE, the algorithm stops if the relative difference in vari-

ance is less than devsm. (default is 0.01)

## Value

## A list with components

cov covariance matrix

mu mean vector

postprob posterior probability

classification classification (0=noise otherwise 1) obtained by rounding postprob

innc list of initial nearest neighbor cleaning results (components are the covariance,

mean, posterior probability and classification)

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## Note

Terms of use: GPL version 2 or newer.

MM: Even though covNNC() is backed by a serious scientific publication, I cannot recommend its use at all.

#### Author(s)

Naisyin Wang <nwang@stat.tamu.edu> and Adrian Raftery <raftery@stat.washington.edu> with contributions from Chris Fraley <fraley@stat.washington.edu>.

covNNC(), then named cov.nnve(), used to be (the only function) in CRAN package **covRobust** (2003), which was archived in 2012.

Martin Maechler allowed ncol(X) == 1, sped up the original code, by reducing the amount of scaling; further, the accuracy was increased (using internal q.dDk()). The original version is available, unexported as robustX:::covNNC1.

#### References

Wang, N. and Raftery, A. (2002) Nearest neighbor variance estimation (NNVE): Robust covariance estimation via nearest neighbor cleaning (with discussion). *Journal of the American Statistical Association* **97**, 994–1019.

See also University of Washington Statistics Technical Report 368 (2000); see at https://stat.uw.edu/research/techreports/

#### See Also

cov.mcd from package MASS; covMcd, and covOGK from package robustbase.

The whole package **rrcov**.

## **Examples**

```
data(iris)
covNNC(iris[-5])

data(hbk, package="robustbase")
hbk.x <- data.matrix(hbk[, 1:3])
covNNC(hbk.x)</pre>
```

L1median

Compute the Multivariate L1-Median aka 'Spatial Median'

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# **Description**

Compute the multivariate  $L_1$ -median m, also called "Spatial Median", i.e., the minimizer of

$$\sum_{i=1}^{n} ||x_i - m||,$$

where 
$$||u|| = \sqrt{\sum_{j=1}^{p} u_j^2}$$
.

As a convex problem, there's always a global minimizer, computable not by a closed formula but rather an iterative search. As the (partial) first derivatives of the objective function is undefined at the data points, the minimization is not entirely trivial.

# Usage

```
L1median(X, m.init = colMedians(X), weights = NULL, method = c("nlm", "HoCrJo", "VardiZhang", optimMethods, nlminbMethods), pscale = apply(abs(centr(X, m.init)), 2, mean, trim = 0.40), tol = 1e-08, maxit = 200, trace = FALSE, zero.tol = 1e-15, ...)
```

# **Arguments**

Χ	numeric matrix of dimension $n \times p$ , say.
m.init	starting value for $m$ ; typically and by default the coordinatewise median.
weights	optional numeric vector of non-negative weights; currently only implemented for method "VardiZhang".
method	character string specifying the computational method, i.e., the algorithm to be used (can be abbreviated).
pscale	numeric p-vector of positive numbers, the coordinate-wise scale (typical size of $\delta m_j$ ), where $m$ is the problem's solution.
tol	positive number specifying the (relative) convergence tolerance.
maxit	positive integer specifying the maximal number of iterations (before the iterations are stopped prematurely if necessary).
trace	an integer specifying the tracing level of the iterations; 0 does no tracing
zero.tol	for method "VardiZhang", a small positive number specifying the tolerance for determining that the iteration is 'exactly' at a data point (which is a singularity).
•••	optional arguments to nlm() or the control (list) arguments of optim(), or nlminb(), respectively.

## **Details**

Currently, we have to refer to the "References" below.

# Value

currently the result *depends* strongly on the method used.

FIXME. This will change considerably.

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#### Author(s)

Martin Maechler. Method "HoCrJo" is mostly based on Kristel Joossens' R function, implementing Hossjer and Croux (1995).

#### References

Hossjer and Croux, C. (1995). Generalizing Univariate Signed Rank Statistics for Testing and Estimating a Multivariate Location Parameter. *Non-parametric Statistics* **4**, 293–308.

Vardi, Y. and Zhang, C.-H. (2000). The multivariate  $L_1$ -median and associated data depth. *Proc. National Academy of Science* **97**(4), 1423–1426.

Fritz, H. and Filzmoser, P. and Croux, C. (2012) A comparison of algorithms for the multivariate L1-median. *Computational Statistics* **27**, 393–410.

Kent, J. T., Er, F. and Constable, P. D. L. (2015) Algorithms for the spatial median;, in K. Nordhausen and S. Taskinen (eds), *Modern Nonparametric, Robust and Multivariate Methods: Festschrift in Honour of Hannu Oja*, Springer International Publishing, chapter 12, pp. 205–224. doi:10.1007/9783319224046\_12

#### See Also

```
median.covMcd
```

CRAN package **pcaPP** added more L1 median methods, re-implementing our R versions in C++, see Fritz et al.(2012) and e.g., 11median\_NLM().

## **Examples**

```
data(stackloss)
L1median(stackloss, method = "HoCrJo")

## Explore all methods:
m <- eval(formals(L1median)$method); allMeths <- m[m != "Brent"]
L1m <- sapply(allMeths, function(meth) L1median(stackloss, method = meth))
## --> with a warning for L-BFGS-B
str(L1m)
pm <- sapply(L1m, function(.) if(is.numeric(.)) . else .$par)
t(pm) # SANN differs a bit; same objective ?</pre>
```

m∨BACON

BACON: Blocked Adaptive Computationally-Efficient Outlier Nominators

#### **Description**

This function performs an outlier identification algorithm to the data in the x array [n x p] and y vector [n] following the lines described by Hadi et al. for their BACON outlier procedure.

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#### Usage

## **Arguments**

x numeric matrix (of dimension [nxp]), not supposed to contain missing values.

collect a multiplication factor c, when init.sel is not "manual", to define m, the size

of the initial basic subset, as  $m\,:=\,c\cdot p,$  in practice, m <-  $\min({\rm p}\star{\rm collect}$  ,

n/2).

m integer in 1:n specifying the *size* of the initial basic subset; used only when

init.sel is not "manual".

alpha determines the cutoff value for the Mahalanobis distances (see details).

init.sel character string, specifying the initial selection mode; implemented modes are:

"Mahalanobis" based on Mahalanobis distances (default); the version V1 of the reference: affine invariant but not robust.

"dUniMedian" based on the distances from the univariate medians; similar to the version V2 of the reference; robust but not affine invariant.

"random" based on a random selection, i.e., reproducible only via set. seed().

"manual" based on manual selection; in this case, a vector man. sel containing the indices of the selected observations must be specified.

"V2" based on the Euclidean norm from the **uni**variate medians; this is the version V2 of the reference; robust but not affine invariant.

"Mahalanobis" and "V2" where proposed by Hadi and the other authors in the reference as versions 'V\_1' and 'V\_2', as well as "manual", while "random" is provided in order to study the behaviour of BACON. Option "dUniMedian" is similar to "V2" and is due to U. Oetliker.

man.sel only when init.sel == "manual", the indices of observations determining the

initial basic subset (and m <- length(man.sel)).</pre>

maxsteps maximal number of iteration steps.

allowSingular logical indicating a solution should be sought also when no matrix of rank p is

found.

verbose logical indicating if messages are printed which trace progress of the algorithm.

## **Details**

Remarks on the tuning parameter alpha: Let  $\chi_p^2$  be a chi-square distributed random variable with p degrees of freedom (p is the number of variables; n is the number of observations). Denote the  $(1-\alpha)$  quantile by  $\chi_p^2(\alpha)$ , e.g.,  $\chi_p^2(0.05)$  is the 0.95 quantile. Following Billor et al. (2000), the cutoff value for the Mahalanobis distances is defined as  $\chi_p(\alpha/n)$  (the square root of  $chi_p^2$ ) times a correction factor c(n,p), n and p, and they use  $\alpha=0.05$ .

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#### Value

#### a list with components

subset	logical vector of length n where the i-th entry is true iff the i-th observation is part of the final selection.
dis	numeric vector of length n with the (Mahalanobis) distances.
cov	$p \times p$ matrix, the corresponding robust estimate of covariance.

## Author(s)

Ueli Oetliker, Swiss Federal Statistical Office, for S-plus 5.1. Port to R, testing etc, by Martin Maechler; Init selection "V2" and correction of default alpha from 0.95 to 0.05, by Tobias Schoch, FHNW Olten, Switzerland.

## References

Billor, N., Hadi, A. S., and Velleman, P. F. (2000). BACON: Blocked Adaptive Computationally-Efficient Outlier Nominators; *Computational Statistics and Data Analysis* **34**, 279–298. doi:10.1016/S01679473(99)001012

#### See Also

covMcd for a high-breakdown (but more computer intensive) method; BACON for a "generalization", notably to *regression*.

```
require(robustbase) # for example data and covMcd():
## simple 2D example :
plot(starsCYG, main = "starsCYG data (n=47)")
B.st <- mvBACON(starsCYG)</pre>
points(starsCYG[ ! B.st$subset,], pch = 4, col = 2, cex = 1.5)
stopifnot(identical(which(!B.st$subset), c(7L,11L,20L,30L,34L)))
## finds the 4 clear outliers (and 1 "borderline");
## it does not find obs. 14 which is an outlier according to covMcd(.)
iniS <- setNames(, eval(formals(mvBACON)$init.sel)) # all initialization methods, incl "random"</pre>
set.seed(123)
Bs.st <- lapply(iniS[iniS != "manual"], function(s)</pre>
                mvBACON(as.matrix(starsCYG), init.sel = s, verbose=FALSE))
ii <- - match("steps", names(Bs.st[[1]]))</pre>
Bs.s1 <- lapply(Bs.st, `[`, ii)
stopifnot(exprs = {
   length(Bs.s1) >= 4
   length(unique(Bs.s1)) == 1 # all 4 methods give the same
})
## Example where "dUniMedian" and "V2" differ :
data(pulpfiber, package="robustbase")
dU.plp <- mvBACON(as.matrix(pulpfiber), init.sel = "dUniMedian")</pre>
V2.plp <- mvBACON(as.matrix(pulpfiber), init.sel = "V2")</pre>
```

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```
(oU <- which(! dU.plp$subset))</pre>
(o2 <- which(! V2.plp$subset))</pre>
stopifnot(setdiff(o2, oU) %in% c(57L,58L,59L,62L))
## and 57, 58, 59, and 62 *are* outliers according to covMcd(.)
## 'coleman' from pkg 'robustbase'
coleman.x <- data.matrix(coleman[, 1:6])</pre>
Cc <- covMcd (coleman.x) # truly robust</pre>
summary(Cc) # -> 6 outliers (1,3,10,12,17,18)
Cb1 <- mvBACON(coleman.x) ##-> subset is all TRUE hmm??
Cb2 <- mvBACON(coleman.x, init.sel = "dUniMedian")</pre>
stopifnot(all.equal(Cb1, Cb2))
## try 20 different random starts:
Cb.r <- lapply(1:20, function(i) { set.seed(i)</pre>
                     mvBACON(coleman.x, init.sel="random", verbose=FALSE) })
nm <- names(Cb.r[[1]]); nm <- nm[nm != "steps"]</pre>
all(eqC <- sapply(Cb.r[-1], function(CC) all.equal(CC[nm], Cb.r[[1]][nm]))) \ \# \ TRUE
## --> BACON always breaks down, i.e., does not see the outliers here
## breaks down even when manually starting with all the non-outliers:
Cb.man <- mvBACON(coleman.x, init.sel = "manual",</pre>
                  man.sel = setdiff(1:20, c(1,3,10,12,17,18)))
which( ! Cb.man$subset) # the outliers according to mvBACON : _none_
```

Orot

Rotation Matrix to Specific Direction

# Description

Construct the  $p \times p$  rotation matrix that rotates the unit vector (1,0,...0), i.e., the  $x_1$ -axis, onto  $(1,1,1,...1)/\sqrt{p}$ , or more generally to  $u/\|u\|$  (u:=unit.image).

#### Usage

```
Qrot(p, transpose = FALSE, unit.image = rep(1, p))
```

# Arguments

p integer; the dimension (of the vectors involved).

transpose logical indicating if the *transposed* matrix is to returned.

 $\verb"unit.image" numeric vector of length $p$ onto which the unit vector should be rotated; defaults$ 

to "the diagonal"  $\propto (1, 1, 1, ..., 1)$ .

#### **Details**

The qr decomposition is used for a Gram-Schmitt basis orthogonalization.

## Value

 $p \times p$  orthogonal matrix which rotates (1,0,...,0) onto a vector proportional to unit.image.

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## Author(s)

Martin Maechler

#### See Also

```
qr, matrix (and vector) multiplication, %*%.
```

# **Examples**

```
Q <- Qrot(6)
zapsmall(crossprod(Q)) # 6 x 6 unity <==> Q'Q = I <==> Q orthogonal

if(require("MASS")) {
   Qt <- Qrot(6, transpose = TRUE)
   stopifnot(all.equal(Qt, t(Q)))
   fractions(Qt ^2) # --> 1/6  1/30  etc, in an almost lower-triagonal matrix
}
```

rbwheel

Multivariate Barrow Wheel Distribution Random Vectors

# **Description**

Generate p-dimensional random vectors according to Stahel's Barrow Wheel Distribution.

## Usage

# **Arguments**

n	integer, specifying the sample size.
р	integer, specifying the dimension (aka number of variables).
frac	numeric, the proportion of outliers. The default, $1/p$ , corresponds to the (asymptotic) breakdown point of M-estimators.
sig1	thickness of the "wheel", (= $\sigma$ (good[,1])), a non-negative numeric.
sig2	thickness of the "axis" (compared to 1).
rGood	function; the generator for "good" observations.
r0ut	function, generating the outlier observations.
U1	p-vector to which $(1,0,\ldots,0)$ is rotated.

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scaleAfter	logical indicating if the matrix is re-scaled <i>after</i> rotation (via scale()) Default TRUE; note that this used to be false by default in the first public version.
scaleBefore	logical indicating if the matrix is re-scaled before rotation (via scale()).
spherize	logical indicating if the matrix is to be "spherized", i.e., rotated and scaled to have empirical covariance $I_p$ . This means that the principal components are used (before rotation).
fullResult	logical indicating if in addition to the $n \times p$ matrix, some intermediate quantities are returned as well.

#### **Details**

....

#### Value

By default (when fullResult is FALSE), an  $n \times p$  matrix of n sample vectors of the p dimensional barrow wheel distribution, with an attribute, n1 specifying the exact number of "good" observations,  $n1 \approx (1-f) \cdot n$ , f =frac.

If fullResult is TRUE, a list with components

```
X the n \times p matrix of above, X = X0 %*% A, where A <- Qrot(p, u = U1), and X0 is the corresponding matrix before rotation, see below.

X0 .......

A the p \times p rotation matrix, see above.

11 the number of "good" observations, see above.

12 the number of "outlying" observations, n2 = n - n1.
```

## Author(s)

Werner Stahel and Martin Maechler

#### References

```
http://stat.ethz.ch/people/maechler/robustness
```

Stahel, W.~A. and Mächler, M. (2009). Comment on "invariant co-ordinate selection", *Journal of the Royal Statistical Society B* **71**, 584–586. doi:10.1111/j.14679868.2009.00706.x

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```
n1 \leftarrow attr(r, "n1") ; pairs(r, col=1+((1:n) > n1))
## for n = 500, you *do* see it :
n <- 500
pairs(r <- rbwheel(n,6))</pre>
## show explicitly
n1 \leftarrow attr(r, "n1") ; pairs(r, col=1+((1:n) > n1))
## but increasing sig2 does help:
pairs(r \leftarrow rbwheel(n,6, sig2 = .2))
## show explicitly
n1 \leftarrow attr(r, "n1") ; pairs(r, col=1+((1:n) > n1))
set.seed(12)
pairs(X <- rbwheel(n, 7, spherize=TRUE))</pre>
colSums(X) # already centered
if(require("ICS") && require("robustbase")) {
  # ICS: Compare M-estimate [Max.Lik. of t_{df} = 2] with high-breakdown:
  stopifnot(require("MASS"))
  X.paM <- ics(X, S1 = cov, S2 = function(.) cov.trob(., nu=2)$cov, stdKurt = FALSE)</pre>
  X.paM. \leftarrow ics(X, S1 = cov, S2 = function(.) tM(., df=2)$V, stdKurt = FALSE)
  X.paR <- ics(X, S1 = cov, S2 = function(.) covMcd(.)$cov, stdKurt = FALSE)</pre>
  plot(X.paM) # not at all clear
  plot(X.paM.)# ditto
  plot(X.paR)# very clear
## Similar such experiments ---> demo(rbwheel_d) and demo(rbwheel_ics)
##
```

reclas

Recursive Robust Median-like Location and Scale

# Description

Calculate an estimate of location, asymptotically equivalent to the median, and an estimate of scale equal to the **MEAN** absolute deviation. Both done recursively.

# Usage

```
reclas(y, b = 0.2, mfn = function(n) 0.1 * n^(-0.25),
    nstart = 30, m0 = median(y0),
    scon=NULL, updateScale = is.null(scon))
```

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## **Arguments**

y numeric vector of i.i.d. data whose location and scale parameters are to be estimated.

b numeric tuning parameter (default value equal to that used by Holst, 1987).

mfn a function of the index of the data which must be positive and and tend to 0 as

the index tends to infinity. The default function is that used by Holst, 1987.

nstart number of starting values: Starting values for the algorithm are formed from the

first nstart values of y. The default value is that used in Cameron and Turner,

1993.

mo value for the initial approximate median; by default, the median of the first

nstart observations.

scon value for the scale parameter s, a function or NULL. When NULL, as by default,

the scale is initialized to the mean of the absolute differences between the first nstart y values and m0. If scon is a function, the initial scale is set to scon(y0, m0), where y0 is the vector of the first nstart y values. Note that scon also

determines the default for updateScale.

updateScale a logical indicating if the scale, initialized from scon should be updated in each

iteration. Otherwise, the the scale is held constant throughout and the algorithm

becomes equivalent to the algorithm of Holst.

#### Value

An S3 "object" of class "reclas"; simply a list with entries

locn the successive recursive estimates of location. The first nstart - 1 of these are

NA.

scale the successive recursive estimates of scale if updateScale is true; otherwise the

constant value used for the scale.

updateScale the same as the function argument.
call the function call, i.e., match.call.

There is a plot method for "reclas", see the examples.

#### Author(s)

```
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```

Extensions by Martin Maechler (scon as function; updateScale, plot()).

#### References

Cameron, Murray A. and Turner, T. Rolf (1993). Recursive location and scale estimators. *Commun. Statist.* — *Theory Meth.* **22**(9) 2503–2515.

Holst, U. (1987). Recursive estimators of location. *Commun. Statist.* — *Theory Meth.* **16** (8) 2201–2226.

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```
set.seed(42)
y <- rt(10000, df = 1.5) # not quite Gaussian ...
z1 <- reclas(y)</pre>
z3 <- reclas(y, scon= 1 ) # correct fixed scale</pre>
z4 <- reclas(y, scon= 100) # wrong fixed scale
z2 <- reclas(y, # a more robust initial scale:</pre>
            scon = function(y0, m0) robustbase::Qn(y0 - m0),
            updateScale = TRUE) # still updated
## Visualizing -- using the plot() method for "reclas":
M \leftarrow median(y) ; yl \leftarrow c(-1,1)* 0.5
OP <- par(mfrow=c(2,2), mar=.1+c(3,3,1,1), mgp=c(1.5, .6, \emptyset))
 plot(z1, M=M, ylim=yl)
 plot(z2, M=M, ylim=yl)
 plot(z3, M=M, ylim=yl)
 plot(z4, M=M, ylim=yl)
par(OP)
```

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